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# A New Cost Effective Estimator in the Presence of Non-response for Two-phase Sampling

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#### Graphical abstract

 $\overline{y}^* = (r_1/n)\overline{y}_{r_11} + (r_2/n)\overline{y}_{k2}$ 

#### Abstract

In In this paper we have proposed a new estimator of population mean in the presence of non-response using information of a single auxiliary variable. We have obtained survey cost for the fixed variance of the proposed estimator and compared it with the cost obtained by Tabasum and Khan (2004) and Hansen Hurwitz (1946). After the comparison we saw that the cost of our proposed estimator is lesser than Tabasum and Khan (2004) and Hansen Hurwitz (1946) estimators.

*Keywords*: Non-response; Hansen Hurwitz estimator; Tabasum and Khan estimator; auxiliary variable; two-phase sampling

### Abstrak

Dalam makalah ini, kami telah mencadangkan satu penganggar baru bagi min populasi dengan kehadiran tak-sambut penggunaan informasi bagi satu pemboleh ubah sokongan. Kami telah mendapatkan kos kaji selidik bagi varian tetap penganggar yang dicadangkan dan dibandingkan dengan kos yang diperoleh oleh penganggar Tabasum dan Khan dan penganggar Hansen Hurwitz. Selepas perbandingan, didapati bahawa kos penganggar yang dicadang adalah yang paling kurang berbanding dengan kos penganggar Tabasum and Khan dan kos penganggar Hansen Hurwitz.

Kata kunci: Tak-sambut; penganggar Hansen Hurwitz; penganggar Tabasum dan Khan; pemboleh ubah sokongan

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# **1.0 INTRODUCTION**

Non-response has been a major problem of almost every sample surveys. The incomplete data create many problems for researcher and this problem cannot be eliminated even by increasing the sample size. The non-response always exists when surveying human populations as people hesitate to respond in surveys. In sensitive issues the non-response rate increases. The pioneer researchers in this area were Hansen and Hurwitz (1946). After that many survey statisticians have suggested methods of estimating population characteristics in the presence of nonresponse. The sub-sampling method has been a popular method in case of non-response. Due to sub sampling the cost survey is increased.

In this paper we have proposed a new estimator for population mean under Two-phase sampling in the presence of non-response. We have also derived mean square error for that proposed estimator and obtain the optimum values of the sample sizes at first phase, second phase and sampling fraction which minimize the survey cost.

We have compared empirically the survey cost of new proposed estimators with the Hansen and Hurwitz (1946) estimator and Tabasum and Khan (2004) estimator. We found that the cost of our proposed estimator is less than the cost obtained by Hansen and Hurwitz (1946) estimator and Tabasum and Khan (2004) estimator.

# **2.0 MATERIALS AND METHODS**

#### 2.1 Two-phase Sampling Scheme

Suppose a simple random sample without replacement (SRSWOR) of size *n* is drawn from a population of size *N*. From the available sample,  $r_1$  units respond to survey variable *Y* and  $r_2$  units do not respond. Corresponding to sample respondents and non–respondents, the population is also divided in same sort of groups containing  $N_1$  and  $N_2$  units. Out of  $r_2$  non–respondents, a

sub-sample of k ( $k=r_2/h$ , h>1) units is drawn and information is obtained from these k units. Hansen and Hurwitz (1946) suggested following estimator of population mean when sub-sampling is used to overcome non-response:

$$\overline{\mathbf{y}}^* = (r_1/n) \overline{\mathbf{y}}_{r_1 1} + (r_2/n) \overline{\mathbf{y}}_{k2}, \qquad (2.1)$$

where  $\overline{y}_1 = r_1^{-1} \sum_{i=1}^{r_i} y_i$  and  $\overline{y}_{k2} = k^{-1} \sum_{i=1}^{k} y_i$  are means of variable of interest. The estimator (2.1) is unbiased with variance:

$$Var(\overline{y}^*) = \lambda_2 S_y^2 + \theta S_{y_2}^2, \qquad (2.2)$$

with 
$$S_y^2 = \sum_{i=1}^N (y_i - \overline{Y})^2 / (N-1)$$
,  $S_{y_2}^2 = \sum_{i=1}^{N_2} (y_i - \overline{Y}_2)^2 / (N_2 - 1)$ ,  
 $\lambda = (1 - f) / n$ ,  $f = n / N$ ,  
 $\theta = W_2 (h-1) / n$ ,  $W_2 = N_2 / N$ ,  $\lambda_1 = n_1^{-1} - N^{-1}$ ,  
 $\lambda_2 = n_2^{-1} - N^{-1}$ ,  $\lambda_3 = n_2^{-1} - n_1^{-1}$   
 $\overline{Y} = N^{-1} \sum_{i=1}^N y_i$  and  $\overline{Y}_2 = N_2^{-1} \sum_{i=1}^{N_2} y_i$ .

## 2.2 Non-response in Two Phase Sampling

The two phase sampling procedure has been effectively used in the presence of non-response to increase the precision of estimates. The two phase sampling procedure in case of nonresponse is described as:

- i) Select a first phase sample of size *n*<sub>1</sub> using SRSWOR and record information on auxiliary variable *X*.
- ii) Select a second phase sample of size  $n_2$  using SRSWOR from first phase sample of size  $n_1$ . The  $r_1$  units respond and  $r_2$  units do not respond. Collect information on study variable *Y* from responding units.
- iii) Select a subsample of size k ( $k=r_2/h$ , h>1) and record information on study variable from these selected units.

Using above two phase sampling procedure, various authors have proposed different estimators of population mean in the presence of non–response. Some notable references are of Cochran (1977), Rao (1986), Naik and Gupta (1991), Tripathi and Khare (1997), Tabasum and Khan (2004, 2006) and Khare and Srivastava (1993, 1995, 2010), Singh and Kumar (2008a, 2008b, 2008c, 2009, 2011).

# 2.3 New Proposed Estimator with Cost Function and Optimum Values Modeling Approach

The proposed estimator for the situation, when non-response occurs in study variable y and auxiliary variable x is

$$t_d = \overline{y}^* - \left(\sqrt{\overline{x}^*} - \sqrt{\overline{x}_1}\right)$$
  
We know that

$$\overline{e}_{y}^{*} = \overline{y}^{*} - \overline{Y} \implies \overline{y}^{*} = \overline{Y} + \overline{e}_{y}^{*}$$

$$\overline{e}_{x}^{*} = \overline{x}^{*} - \overline{X} \implies \overline{x}^{*} = \overline{X} + \overline{e}_{x}^{*}$$

$$\overline{\mathbf{e}}_{\mathbf{x}_1} = \overline{\mathbf{x}}_1 - \mathbf{X} \quad \Rightarrow \quad \overline{\mathbf{x}}_1 = \mathbf{X} + \overline{\mathbf{e}}_{\mathbf{x}_1}$$

Putting the values of  $\overline{\mathbf{e}}_{y}^{*}$ ,  $\overline{\mathbf{e}}_{x}^{*}$ , and  $\overline{\mathbf{e}}_{x_{1}}$  in (1) we get  $t_{d} = (\overline{\mathbf{Y}} + \overline{\mathbf{e}}_{y}^{*}) - (\sqrt{\overline{\mathbf{X}} + \overline{\mathbf{e}}_{x}^{*}} - \sqrt{\overline{\mathbf{X}} + \overline{\mathbf{e}}_{x_{1}}})$ 

$$\begin{split} t_d &= \left(\overline{\mathbf{Y}} + \overline{\mathbf{e}}_y^*\right) - \left(\sqrt{\overline{\mathbf{X}}} + \overline{\mathbf{e}}_x^* - \sqrt{\overline{\mathbf{X}}} + \overline{\mathbf{e}}_{\mathbf{x}_1}\right)\\ \text{or}\\ t_d - \overline{\mathbf{Y}} &= \overline{\mathbf{e}}_y^* - \frac{1}{2\sqrt{\overline{\mathbf{X}}}} \left(\overline{\mathbf{e}}_x^* - \overline{\mathbf{e}}_{\mathbf{x}_1}\right) \end{split}$$

Taking Square and apply Expectation on both sides, we have  $MSE(t_d) \approx E(\overline{e}_y^{*2}) + \frac{1}{4\overline{X}} \Big[ E(\overline{e}_x^{*2}) + E(\overline{e}_{x_1}^2) - 2E(\overline{e}_x^*\overline{e}_{x_1}) \Big] - \frac{1}{\sqrt{\overline{X}}} \Big[ E(\overline{e}_y^*\overline{e}_{x_1}^*) - E(\overline{e}_y^*\overline{e}_{x_1}) \Big]$ 

We know that

$$\begin{split} & E\left(\overline{e}_{y}^{*2}\right) = \lambda_{2}S_{y}^{2} + \theta S_{y_{2}}^{2} , E\left(\overline{e}_{y}^{*}\overline{e}_{x_{1}}\right) = \lambda_{1}S_{xy} \\ & E\left(\overline{e}_{x}^{*2}\right) = \lambda_{2}S_{x}^{2} + \theta S_{x_{2}}^{2} , E\left(\overline{e}_{x}^{*}\overline{e}_{y}^{*}\right) = \lambda_{2}S_{xy} + \theta S_{xy2} \\ & E\left(\overline{e}_{x_{1}}^{2}\right) = \lambda_{1}S_{x}^{2} , E\left(\overline{e}_{x_{1}}\overline{e}_{x}^{*}\right) = \lambda_{1}S_{x}^{2} \text{ and } R = \frac{\overline{Y}}{\overline{X}} \\ & \text{We can write} \\ & \text{Var}(T_{R_{1d}}) = \lambda_{2}S_{y}^{2} + \theta S_{y_{2}}^{2} + \frac{1}{4\overline{X}}\left(\lambda_{2}S_{x}^{2} + \theta S_{x_{2}}^{2} + \lambda_{1}S_{x}^{2} - 2\lambda_{1}S_{x}^{2}\right) - \frac{1}{\sqrt{\overline{X}}}\left[\left(\lambda_{2}S_{xy} + \theta S_{xy2}^{2} - \lambda_{1}S_{xy}\right)\right] \end{split}$$

or

$$\begin{split} \textit{MSE}(t_{a}) &\approx \lambda_{3} \bigg( S_{y}^{2} + \frac{S_{x}^{2}}{4\overline{X}} - \frac{S_{yy}}{\sqrt{\overline{X}}} \bigg) + \lambda_{3} S_{y}^{2} + \theta \bigg( S_{y_{2}}^{2} + \frac{S_{x_{2}}}{4\overline{X}} - \frac{S_{yy_{2}}}{\sqrt{\overline{X}}} \bigg) \\ \textit{Let us consider a cost function} \end{split}$$

 $C = c_1 n_1 + c_2 n_2 + c_3 r_1 + c_4 k$ 

Where

 $c_1$  = The unit cost associated with first phase sample,  $n_1$ 

 $c_2$  = The cost of first attempt on Y with second phase sample,  $n_2$ 

 $c_3 =$  The unit cost for processing the respondent data on Y at the first attempt in  $r_1$ 

 $c_4 =$  The unit cost associated with the sub-sample k of  $r_2$ 

Since the value of  $r_1$  and k is not known until the first attempt is made, so the expected cost will be used in planning survey. The expected value of  $r_1$  and k are  $W_1n_2$  and  $\frac{W_2n_2}{h}$ . Thus the

expected cost is given by

$$E(C) = C^* = c_1 n_1 + \left(c_2 + c_3 W_1 + \frac{c_4 W_2}{h}\right) n_2$$

To determine the optimum values of h,  $n_2$ , and  $n_1$  that minimize the cost for a fixed variance  $V_0$  we consider the function

$$\phi = \mathbf{C}^* + \lambda \left\{ MSE(t_d) - \mathbf{V}_o \right\}$$

$$\phi = c_1 n_1 + \left( c_2 + c_3 W_1 + \frac{c_4 W_2}{h} \right) n_2 + \lambda \left\{ \left( \frac{1}{n_1} - \frac{1}{N} \right) S_y^2 + \left( \frac{1}{n_2} - \frac{1}{n_1} \right) S_r^2 + \left( \frac{W_2(h-1)}{n_2} \right) S_{2r}^2 - V_o \right\}$$

Where

$$\begin{split} S_r^2 &= S_y^2 + \frac{S_x^2}{4\overline{X}} - \frac{S_{xy}}{\sqrt{\overline{X}}} \\ S_{r_2}^2 &= S_{y_2}^2 + \frac{S_{x_2}^2}{4\overline{X}} - \frac{S_{xy2}}{\sqrt{\overline{X}}} \end{split}$$

Where  $\lambda$  is Lagrange's multiplier.

Using Lagrange's multiplier technique the optimum values  $h,\,n_2$  and  $n_1$  are

$$\phi = c_1 n_1 + \left(c_2 + c_3 W_1 + \frac{c_4 W_2}{h}\right) n_2 + \lambda \left\{ \left(\frac{1}{n_1} - \frac{1}{N}\right) S_y^2 + \left(\frac{1}{n_2} - \frac{1}{n_1}\right) S_r^2 + \left(\frac{W_2(h-1)}{n_2}\right) S_{r_2}^2 - V_o \right\}$$

For optimum value of  $\lambda,$  differentiate w.r.t.  $\lambda$  and equate to zero.

$$\left(\frac{1}{n_1}\right)\left(S_y^2 - S_r^2\right) + \frac{1}{n^2}\left[S_r^2 + w_2(h-1)S_r^2\right] = V_0 + \frac{S_y^2}{N}$$
(2.3)

Now we differentiate w.r.t. h and equate to zero.

Muhammad Ismail, Muhammad Hanif & Muhammad Qaiser / Jurnal Teknologi (Sciences & Engineering) 63:2 (2013), 1-4

$$h^{2} = \frac{n_{2}^{2}c_{4}}{\lambda S_{*}^{2}}$$
(2.4)

Now we differentiate w.r.t.  $n_2$ , we get

$$n_{2}^{2} = \frac{\lambda \left(S_{r}^{2} + W_{2}(h-1)S_{r_{2}}^{2}\right)}{c_{2} + c_{3}W_{1} + \frac{c_{4}W_{2}}{h}}$$

By putting the value of  $n_2^2$  in (2.4), we get the value of "h"

$$h = \sqrt{\frac{c_4 \left(S_r^2 - W_2 S_{r_2}^2\right)}{S_{r_2}^2 \left(c_2 + c_3 W_1\right)}}$$

Differentiate w.r.t. n<sub>1</sub>, we get

$$\mathbf{n}_1 = \sqrt{\frac{\lambda}{c_1}} \Big( \mathbf{S}_y^2 - \mathbf{S}_r^2 \Big)$$

Putting the value of  $n_1$  in equation(2.3)

$$\sqrt{\lambda} = \frac{\sqrt{c_1 \left(S_y^2 - S_r^2\right)} + \left(\sqrt{c_2 + c_3 W_1 + \frac{c_4 W_2}{h}}\right) \left(\sqrt{S_r^2 + W_2 (h-1) S_{r_2}^2}\right)}{V_o + \frac{S_y^2}{N}}$$

Putting the value of  $\sqrt{\lambda}$  in equation (2.3)

$$n_{1} = \frac{\left[\sqrt{c_{1}\left(S_{y}^{2} - S_{r}^{2}\right)} + \left(\sqrt{c_{2} + c_{3}W_{1} + \frac{c_{4}W_{2}}{h}}\right)\left(\sqrt{S_{r}^{2} + W_{2}(h-1)S_{r_{2}}^{2}}\right)\right]\sqrt{S_{y}^{2} - S_{r}^{2}}}{\left(V_{o} + \frac{S_{y}^{2}}{N}\right)\sqrt{c_{1}}}$$

We have

$$n^{}_{2} = \sqrt{\frac{\lambda \Big(S^{2}_{r} \ + \ W^{}_{2}(h-l)S^{2}_{r^{}_{r}}\Big)}{c^{}_{2} \ + \ c^{}_{3}W^{}_{1} \ + \ \frac{c^{}_{4}W^{}_{2}}{h}}}$$

Replace the value of  $\boldsymbol{\lambda}$  in above

$$n_{2} = \frac{\sqrt{S_{r}^{2} + W_{2}(h-1)S_{r_{2}}^{2}} \left[ \sqrt{c_{1}\left(S_{y}^{2} - S_{r}^{2}\right)} + \left(\sqrt{c_{2} + c_{3}W_{1} + \frac{c_{4}W_{2}}{h}}\right) \left(\sqrt{S_{r}^{2} + W_{2}(h-1)S_{r_{2}}^{2}}\right) - \frac{\left(V_{o} + \frac{S_{y}^{2}}{N}\right) \sqrt{c_{2} + c_{3}W_{1} + \frac{c_{4}W_{2}}{h}}}{\left(V_{o} + \frac{S_{y}^{2}}{N}\right) \sqrt{c_{2} + c_{3}W_{1} + \frac{c_{4}W_{2}}{h}}}$$

2.4 Cost Function and Optimum and Values in Hansen Hurwitz Estimator

The variance of the Hansen Hurwitz Estimator  $\overline{y}^*$  is

$$Var\left(\overline{y}^*\right) = \lambda_2 S_y^2 + \theta S_{y_2}^2$$

The expected cost function is given by this

$$C_1^* = \left(c_2 + c_3 W_1 + \frac{c_4 W_2}{h}\right) n_2$$

To determine the optimum values of h, and  $n_2$  that minimize the cost for a fixed variance  $V_0\,we$  consider the function

$$\begin{split} \phi &= \mathbf{C}_1^* + \lambda \Big\{ \mathbf{Var} \Big( \overline{y}^* \Big) - \mathbf{V}_{o} \Big\} \\ \phi &= \Big( \mathbf{c}_2 + \mathbf{c}_3 \mathbf{W}_1 + \frac{\mathbf{c}_4 \mathbf{W}_2}{\mathbf{h}} \Big) \mathbf{n}_2 + \lambda \Big\{ \Big( \frac{1}{\mathbf{n}_2} - \frac{1}{\mathbf{N}} \Big) \mathbf{S}_y^2 + \Big( \frac{\mathbf{W}_2(\mathbf{h} - \mathbf{l})}{\mathbf{n}_2} \Big) \mathbf{S}_{y_2}^2 - \mathbf{V}_{o} \Big\} \end{split}$$

Where  $\lambda$  is Lagrange's multiplier.

Using Lagrange's multiplier technique the optimum values h,  $n_2$  and  $n_1$  are

$$\begin{split} h_{_{oHH}} = & \sqrt{\frac{c_{_{4}} \left(S_{_{y}}^{2} - W_{_{2}}S_{_{y_{2}}}^{2}\right)}{S_{_{y_{2}}^{2}}^{2} \left(c_{_{2}} + c_{_{3}}W_{_{1}}\right)}}, \\ n_{_{2HH}} = \frac{S_{_{y}}^{2} + W_{_{2}}(h-1)S_{_{y_{2}}}^{2}}{\left(V_{_{o}} + \frac{S_{_{y}}^{2}}{N}\right)} \end{split}$$

# 2.5 Cost Function and Optimum Values in Tabasum and Khan (2004)

Tabasum and Khan (2004) defined the double sampling ratio estimator as

$$t_{tk} = \overline{y}^* \left( \overline{x}_1 / \overline{x}^* \right)$$

(4.4) The approximate mean square error  $t_{tk}$  given by

$$MSE(t_{ik}) \approx \left(\frac{1}{n_1} - \frac{1}{N}\right) S_y^2 + \left(\frac{1}{n_2} - \frac{1}{n_1}\right) S_r^2 + \left(\frac{W_2(h-1)}{n_2}\right) S_2^2$$
  
The expected cost function is given by

$$\mathbf{C}_{2}^{*} = \mathbf{c}_{1}\mathbf{n}_{1} + \left(\mathbf{c}_{2} + \mathbf{c}_{3}\mathbf{W}_{1} + \frac{\mathbf{c}_{4}\mathbf{W}_{2}}{\mathbf{h}}\right)\mathbf{n}_{2}$$

To determine the optimum value hotk

$$\phi = \mathbf{C}_{2}^{*} + \lambda \left\{ MSE(t_{tk}) - \mathbf{V}_{o} \right\}$$
  
 
$$\phi = c_{1}n_{1} + \left( c_{2} + c_{3}\mathbf{W}_{1} + \frac{c_{4}\mathbf{W}_{2}}{h} \right)n_{2} + \lambda \left\{ \left( \frac{1}{n_{1}} - \frac{1}{N} \right) \mathbf{S}_{y}^{2} + \left( \frac{1}{n_{2}} - \frac{1}{n_{1}} \right) \mathbf{S}_{r}^{2} + \left( \frac{\mathbf{W}_{2}(\mathbf{h} - \mathbf{l})}{n_{2}} \right) \mathbf{S}_{2r}^{2} - \mathbf{V}_{o} \right\}$$

Where

$$\begin{split} S_r^2 &= S_y^2 + R^2 S_x^2 \, - \, 2RS_{xy} \\ S_{r_2}^2 &= S_{y_2}^2 + R^2 S_{x_2}^2 \, - \, 2RS_{xy_2} \end{split}$$

Where  $\lambda$  is Lagrange's multiplier.

Using Lagrange's multiplier technique the optimum values  $h,\,n_2$  and  $n_1$  are

$$\begin{split} \phi &= c_{1}n_{1} + \left(c_{2} + c_{3}W_{1} + \frac{c_{4}W_{2}}{h}\right)n_{2} + \lambda \left\{ \left(\frac{1}{n_{1}} - \frac{1}{N}\right)S_{y}^{2} + \left(\frac{1}{n_{2}} - \frac{1}{n_{1}}\right)S_{r}^{2} + \left(\frac{W_{2}(h-l)}{n_{2}}\right)S_{r_{2}}^{2} - V_{o} \right\} \\ h_{oTK} &= \sqrt{\frac{c_{4}\left(S_{r}^{2} - W_{2}S_{r_{2}}^{2}\right)}{S_{r_{2}}^{2}\left(c_{2} + c_{3}W_{1}\right)}}, \\ n_{2TK} &= \frac{\sqrt{S_{r}^{2} + W_{2}(h-l)S_{r_{2}}^{2}}\left[\sqrt{c_{1}(S_{y}^{2} - S_{r}^{2})} + \left(\sqrt{c_{2} + c_{3}W_{1} + \frac{c_{4}W_{2}}{h}}\right)\left(\sqrt{S_{r}^{2} + W_{2}(h-l)S_{r_{2}}^{2}}\right)\right]}{\left(V_{o} + \frac{S_{y}^{2}}{N}\right)\sqrt{c_{2} + c_{3}W_{1} + \frac{c_{4}W_{2}}{h}}} \\ n_{1TK} &= \frac{\left[\sqrt{c_{1}(S_{y}^{2} - S_{r}^{2})} + \left(\sqrt{c_{2} + c_{3}W_{1} + \frac{c_{4}W_{2}}{h}}\right)\left(\sqrt{S_{r}^{2} + W_{2}(h-l)S_{r_{2}}^{2}}\right)\right]\sqrt{S_{y}^{2} - S_{r}^{2}}}{\left(V_{o} + \frac{S_{y}^{2}}{N}\right)\sqrt{c_{1}}} \end{split}$$

## 2.6 Empirical Comparison of the Estimators

The expected cost  $\mathbf{C}^*$  for our proposed estimator  $t_d$  and expected cost  $\mathbf{C}_1^*$  for Hansen Hurwitz estimator  $\overline{\mathbf{y}}^*$  and  $\mathbf{C}_2^*$  is

the expected cost for  $t_{ik}$  are compared by using population of Tabasum and Khan (2004) paper. The parameters of the population are

$$\begin{split} & \mathbf{N}_1 = 500, \, \mathbf{N}_2 = 150, \, \mathbf{R} = 1.48, \, \rho_1 = 0.81, \, \mathbf{S}_x^2 = 350.54 \,, \\ & \mathbf{S}_y^2 = 1213.82, \, \mathbf{S}_{xy} = 530.07 \,, \, \mathbf{S}_{x_2}^2 = 150.04 \,, \\ & \mathbf{S}_{y_2}^2 = 610.67 \,, \, \mathbf{S}_{xy_2} = 253.68 \,, \beta_1 = 1.69, \, \beta_2 = 1.69, \, \rho_2 = 0.83, \\ & \overline{\mathbf{X}} = 500 \end{split}$$

Table 1 Expected cost for fixed variance

W <sub>1</sub>	W <sub>2</sub>	c <sub>1</sub>	<b>c</b> <sub>2</sub>	c <sub>3</sub>	c4	For Fixed Variance V <sub>o</sub> = 5.41		
						Expected cost $C^*$	Expected cost $C_1^*$	Expected cost C <sup>*</sup> <sub>2</sub>
0.7	0.3	0.1	0.5	1	2	91	265	160
		0.2	0.6	1.4	3	135	361	241
		0.3	0.8	1.6	4	177	448	317
		0.4	0.9	1.9	5	219	531	390

# **3.0 CONCLUSION**

It is observe that the expected cost  $C^*$  for our proposed estimator  $t_d$  is lesser than and expected cost  $C_1^*$  for Hansen Hurwitz

estimator  $\overline{y}^*$  and  $C_2^*$  is the expected cost for  $t_{tk}$  Tabasum and Khan (2004).

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